

Source of the Data and Accuracy of the Estimates for the October 2017 CPS Microdata File

SOURCE OF THE DATA

The data in this microdata file are from the October 2017 Current Population Survey (CPS). The U.S. Census Bureau conducts the CPS every month, although this file has only October data. The October survey uses two sets of questions, the basic CPS and a set of supplemental questions. The CPS, sponsored jointly by the Census Bureau and the U.S. Bureau of Labor Statistics, is the country's primary source of labor force statistics for the civilian noninstitutionalized population. The Census Bureau and the National Center for Educational Statistics also jointly sponsor the supplemental questions for October.

Basic CPS. The monthly CPS collects primarily labor force data about the civilian noninstitutionalized population living in the United States. The institutionalized population, which is excluded from the population universe, is composed primarily of the population in correctional institutions and nursing homes (98 percent of the 4.0 million institutionalized people in Census 2010). Interviewers ask questions concerning labor force participation about each member 15 years old and over in sample households. Typically, the week containing the nineteenth of the month is the interview week. The week containing the twelfth is the reference week (i.e., the week about which the labor force questions are asked).

The CPS uses a multistage probability sample based on the results of the decennial census, with coverage in all 50 states and the District of Columbia. The sample is continually updated to account for new residential construction. When files from the most recent decennial census become available, the Census Bureau gradually introduces a new sample design for the CPS.

Every ten years, the CPS first stage sample is redesigned¹ reflecting changes based on the most recent decennial census. In the first stage of the sampling process, primary sampling units (PSUs)² were selected for sample. In the 2000 design, the United States was divided into 2,025 PSUs. These were then grouped into 824 strata and one PSU was selected for sample from each stratum. In the 2010 sample design, the United States was divided into 1,987 PSUs. These PSUs were then grouped into 852 strata. Within each stratum, a single PSU was chosen for the sample, with its probability of selection proportional to its population as of the most recent decennial census. In the case of strata consisting of only one PSU, the PSU was chosen with certainty.

In April 2014, the Census Bureau began phasing out the 2000 sample and replaced it with the 2010 sample, creating a mixed sampling frame. Two simultaneous changes occurred during this

¹ For detailed information on the 2000 sample redesign, please see reference [1].

² The PSUs correspond to substate areas (i.e., counties or groups of counties) that are geographically contiguous.

phase-in period. First, within the PSUs selected for both the 2000 and 2010 designs, sample households from the 2010 design gradually replaced sample households from the 2000 design. Second, new PSUs selected for only the 2010 design gradually replaced outgoing PSUs selected for only the 2000 design. By July 2015, the new 2010 sample design was completely implemented and the sample came entirely from the 2010 redesigned sample.

Approximately 72,000 housing units were selected for sample from the sampling frame in October. Based on eligibility criteria, 10 percent of these housing units were sent directly to computer-assisted telephone interviewing (CATI). The remaining units were assigned to interviewers for computer-assisted personal interviewing (CAPI).³ Of all housing units in sample, about 61,000 were determined to be eligible for interview. Interviewers obtained interviews at about 53,000 of these units. Noninterviews occur when the occupants are not found at home after repeated calls or are unavailable for some other reason.

October 2017 Supplement. In October 2017, in addition to the basic CPS questions, interviewers asked supplementary questions of household members three years old and over on school enrollment.

Estimation Procedure. This survey's estimation procedure adjusts weighted sample results to agree with independently derived population estimates of the civilian noninstitutionalized population of the United States and each state (including the District of Columbia). These population estimates, used as controls for the CPS, are prepared monthly to agree with the most current set of population estimates that are released as part of the Census Bureau's population estimates and projections program.

The population controls for the nation are distributed by demographic characteristics in two ways:

- Age, sex, and race (White alone, Black alone, and all other groups combined).
- Age, sex, and Hispanic origin.

The population controls for the states are distributed by race (Black alone and all other race groups combined), age (0-15, 16-44, and 45 and over), and sex.

The independent estimates by age, sex, race, and Hispanic origin, and for states by selected age groups and broad race categories, are developed using the basic demographic accounting formula whereby the population from the 2010 Census data is updated using data on the components of population change (births, deaths, and net international migration) with net internal migration as an additional component in the state population estimates.

³ For further information on CATI and CAPI and the eligibility criteria, please see reference [2].

The net international migration component of the population estimates includes:

- Net international migration of the foreign born;
- Net migration between the United States and Puerto Rico;
- Net migration of natives to and from the United States; and
- Net movement of the Armed Forces population to and from the United States.

Because the latest available information on these components lags the survey date, it is necessary to make short-term projections of these components to develop the estimate for the survey date.

ACCURACY OF THE ESTIMATES

A sample survey estimate has two types of error: sampling and nonsampling. The accuracy of an estimate depends on both types of error. The nature of the sampling error is known given the survey design; the full extent of the nonsampling error is unknown.

Sampling Error. Since the CPS estimates come from a sample, they may differ from figures from an enumeration of the entire population using the same questionnaires, instructions, and enumerators. For a given estimator, the difference between an estimate based on a sample and the estimate that would result if the sample were to include the entire population is known as sampling error. Standard errors, as calculated by methods described in “Standard Errors and Their Use,” are primarily measures of the magnitude of sampling error. However, they may include some nonsampling error.

Nonsampling Error. For a given estimator, the difference between the estimate that would result if the sample were to include the entire population and the true population value being estimated is known as nonsampling error. There are several sources of nonsampling error that may occur during the development or execution of the survey. It can occur because of circumstances created by the interviewer, the respondent, the survey instrument, or the way the data are collected and processed. For example, errors could occur because:

- The interviewer records the wrong answer, the respondent provides incorrect information, the respondent estimates the requested information, or an unclear survey question is misunderstood by the respondent (measurement error).
- Some individuals who should have been included in the survey frame were missed (coverage error).
- Responses are not collected from all those in the sample or the respondent is unwilling to provide information (nonresponse error).
- Values are estimated imprecisely for missing data (imputation error).
- Forms may be lost, data may be incorrectly keyed, coded, or recoded, etc. (processing error).

To minimize these errors, the Census Bureau applies quality control procedures during all stages of the production process including the design of the survey, the wording of questions, the review of the work of interviewers and coders, and the statistical review of reports.

Two types of nonsampling error that can be examined to a limited extent are nonresponse and undercoverage.

Nonresponse. The effect of nonresponse cannot be measured directly, but one indication of its potential effect is the nonresponse rate. For the October 2017 basic CPS, the household-level nonresponse rate was 13.8 percent. The person-level nonresponse rate for the School Enrollment supplement was an additional 9.9 percent.

Since the basic CPS nonresponse rate is a household-level rate and the School Enrollment supplement nonresponse rate is a person-level rate, we cannot combine these rates to derive an overall nonresponse rate. Nonresponding households may have fewer persons than interviewed ones, so combining these rates may lead to an overestimate of the true overall nonresponse rate for persons for the School Enrollment supplement.

In accordance with Census Bureau and Office of Management and Budget Quality Standards, the Census Bureau will conduct a nonresponse bias analysis to assess nonresponse bias in the School Enrollment.

Sufficient Partial Interview. A sufficient partial interview is an incomplete interview in which the household or person answered enough of the questionnaire for the supplement sponsor to consider the interview complete. The remaining supplement questions may have been edited or imputed to fill in missing values. Insufficient partial interviews are considered to be nonrespondents. Refer to the supplement overview attachment in the technical documentation for the specific questions deemed critical by the sponsor as necessary to be answered in order to be considered a sufficient partial interview.

As part of the nonsampling error analysis, the item response rates, item refusal rates, and edits are reviewed. For the School Enrollment supplement, the item refusal rates range from 0.0 percent to 1.7 percent. The item allocation rates range from 4.5 percent to 20.6 percent

Coverage. The concept of coverage in the survey sampling process is the extent to which the total population that could be selected for sample “covers” the survey’s target population. Missed housing units and missed people within sample households create undercoverage in the CPS. Overall CPS undercoverage for October 2017 is estimated to be about 11 percent. CPS coverage varies with age, sex, and race. Generally, coverage is larger for females than for males and larger for non-Blacks than for Blacks. This differential coverage is a general problem for most household-based surveys.

The CPS weighting procedure partially corrects for bias from undercoverage, but biases may still be present when people who are missed by the survey differ from those interviewed in ways other than age, race, sex, Hispanic origin, and state of residence. How this weighting procedure affects other variables in the survey is not precisely known. All of these considerations affect comparisons across different surveys or data sources.

A common measure of survey coverage is the coverage ratio, calculated as the estimated population before poststratification divided by the independent population control. Table 1 shows October 2017 CPS coverage ratios by age and sex for certain race and Hispanic groups. The CPS coverage ratios can exhibit some variability from month to month.

Age group	<u>Total</u>		<u>White only</u>		<u>Black only</u>		<u>Residual race</u>		<u>Hispanic</u>		
	All people	Male	Female	Male	Female	Male	Female	Male	Female	Male	Female
0-15	0.87	0.87	0.87	0.90	0.92	0.76	0.74	0.85	0.79	0.79	0.79
16-19	0.88	0.92	0.84	0.96	0.87	0.78	0.72	0.84	0.82	0.94	0.83
20-24	0.76	0.76	0.77	0.79	0.79	0.67	0.64	0.69	0.78	0.78	0.78
25-34	0.83	0.81	0.85	0.84	0.87	0.66	0.76	0.79	0.80	0.72	0.81
35-44	0.89	0.87	0.91	0.90	0.95	0.73	0.80	0.83	0.84	0.81	0.87
45-54	0.92	0.89	0.94	0.91	0.95	0.82	0.89	0.83	0.91	0.81	0.92
55-64	0.93	0.92	0.95	0.93	0.97	0.85	0.87	0.91	0.87	0.85	0.88
65+	0.96	0.98	0.95	0.98	0.96	0.96	0.97	0.88	0.88	0.84	0.86
15+	0.90	0.88	0.90	0.91	0.93	0.78	0.82	0.83	0.84	0.81	0.85
0+	0.89	0.88	0.90	0.91	0.92	0.77	0.81	0.83	0.83	0.80	0.83

- Notes: (1) The Residual race group includes cases indicating a single race other than White or Black, and cases indicating two or more races.
(2) Hispanics may be any race. For a more detailed discussion on the use of parameters for race and ethnicity, please see the “Generalized Variance Parameters” section.

Comparability of Data. Data obtained from the CPS and other sources are not entirely comparable. This results from differences in interviewer training and experience and in differing survey processes. This is an example of nonsampling variability not reflected in the standard errors. Therefore, caution should be used when comparing results from different sources.

Data users should be careful when comparing the data from this microdata file, which reflects Census 2010-based controls, with microdata files from January 2003 through December 2011, which reflect 2000 census-based controls. Ideally, the same population controls should be used when comparing any estimates. In reality, the use of the same population controls is not practical when comparing trend data over a period of 10 to 20 years. Thus, when it is necessary to combine or compare data based on different controls or different designs, data users should be aware that changes in weighting controls or weighting procedures can create small differences between estimates. See the discussion following for information on comparing estimates derived from different controls or different sample designs.

Microdata files from previous years reflect the latest available census-based controls. Although the most recent change in population controls had relatively little impact on summary measures such as averages, medians, and percentage distributions, it did have a significant impact on levels. For example, use of Census 2010-based controls results in about a 0.2 percent increase from the 2000 census-based controls in the civilian noninstitutionalized population and in the number of families and households. Thus, estimates of levels for data collected in 2012 and later years will differ from those for earlier years by more than what could be attributed to actual changes in the population. These differences could be disproportionately greater for certain population subgroups than for the total population.

Users should also exercise caution because of changes caused by the phase-in of the Census 2010 files (see “Basic CPS”).⁴ During this time period, CPS data were collected from sample designs based on different censuses. Two features of the new CPS design have the potential of affecting published estimates: (1) the temporary disruption of the rotation pattern from August 2014 through June 2015 for a comparatively small portion of the sample and (2) the change in sample areas. Most of the known effect on estimates during and after the sample redesign will be the result of changing from 2000 to 2010 geographic definitions. Research has shown that the national-level estimates of the metropolitan and nonmetropolitan populations should not change appreciably because of the new sample design. However, users should still exercise caution when comparing metropolitan and nonmetropolitan estimates across years with a design change, especially at the state level.

Caution should also be used when comparing Hispanic estimates over time. No independent population control totals for people of Hispanic origin were used before 1985.

A Nonsampling Error Warning. Since the full extent of the nonsampling error is unknown, one should be particularly careful when interpreting results based on small differences between estimates. The Census Bureau recommends that data users incorporate information about nonsampling errors into their analyses, as nonsampling error could impact the conclusions drawn from the results. Caution should also be used when interpreting results based on a relatively small number of cases. Summary measures (such as medians and percentage distributions) probably do not reveal useful information when computed on a subpopulation smaller than 75,000.

For additional information on nonsampling error, including the possible impact on CPS data, when known, refer to references [2] and [3].

Standard Errors and Their Use. The sample estimate and its standard error enable one to construct a confidence interval. A confidence interval is a range about a given estimate that has a specified probability of containing the average result of all possible samples. For example, if all possible samples were surveyed under essentially the same general conditions and using the same sample design, and if an estimate and its standard error were calculated from each sample, then approximately 90 percent of the intervals from 1.645 standard errors below the estimate to 1.645 standard errors above the estimate would include the average result of all possible samples.

⁴ The phase-in process using the 2010 Census files began April 2014.

A particular confidence interval may or may not contain the average estimate derived from all possible samples, but one can say with specified confidence that the interval includes the average estimate calculated from all possible samples.

Standard errors may also be used to perform hypothesis testing, a procedure for distinguishing between population parameters using sample estimates. The most common type of hypothesis is that the population parameters are different. An example of this would be comparing the percentage of men who were part-time workers to the percentage of women who were part-time workers.

Tests may be performed at various levels of significance. A significance level is the probability of concluding that the characteristics are different when, in fact, they are the same. For example, to conclude that two characteristics are different at the 0.10 level of significance, the absolute value of the estimated difference between characteristics must be greater than or equal to 1.645 times the standard error of the difference.

The Census Bureau uses 90-percent confidence intervals and 0.10 levels of significance to determine statistical validity. Consult standard statistical textbooks for alternative criteria.

Estimating Standard Errors. The Census Bureau uses replication methods to estimate the standard errors of CPS estimates. These methods primarily measure the magnitude of sampling error. However, they do measure some effects of nonsampling error as well. They do not measure systematic biases in the data associated with nonsampling error. Bias is the average over all possible samples of the differences between the sample estimates and the true value.

There are two ways to calculate standard errors for the CPS microdata file on School Enrollment. They are:

- Direct estimates created from replicate weighting methods;
- Generalized variance estimates created from generalized variance function parameters a and b .

While replicate weighting methods provide the most accurate variance estimates, this approach requires more computing resources and more expertise on the part of the user. The Generalized Variance Function (GVF) parameters provide a method of balancing accuracy with resource usage as well as a smoothing effect on standard error estimates across time. For more information on calculating direct estimates, see reference [4]. For more information on GVF estimates, refer to the “Generalized Variance Parameters” section.

Generalized Variance Parameters. While it is possible to compute and present an estimate of the standard error based on the survey data for each estimate in a report, there are a number of reasons why this is not done. A presentation of the individual standard errors would be of limited use, since one could not possibly predict all of the combinations of results that may be of interest to data users. Additionally, data users have access to CPS microdata files, and it is

impossible to compute in advance the standard error for every estimate one might obtain from those data sets. Moreover, variance estimates are based on sample data and have variances of their own. Therefore, some methods of stabilizing these estimates of variance, for example, by generalizing or averaging over time, may be used to improve their reliability.

Experience has shown that certain groups of estimates have similar relationships between their variances and expected values. Modeling or generalizing may provide more stable variance estimates by taking advantage of these similarities. The GVF is a simple model that expresses the variance as a function of the expected value of the survey estimate. The parameters of the GVF are estimated using direct replicate variances. These GVF parameters provide a relatively easy method to obtain approximate standard errors for numerous characteristics. In this source and accuracy statement, Table 4 provides the GVF parameters for labor force estimates, and Table 5 provides GVF parameters for characteristics from the October 2017 supplement. Table 6 provides factors and population controls to derive regional parameters.

The basic CPS questionnaire records the race and ethnicity of each respondent. With respect to race, a respondent can be White, Black, Asian, American Indian and Alaskan Native (AIAN), Native Hawaiian and Other Pacific Islander (NHOPI), or combinations of two or more of the preceding. A respondent’s ethnicity can be Hispanic or non-Hispanic, regardless of race.

The GVF parameters to use in computing standard errors are dependent upon the race/ethnicity group of interest. The following table summarizes the relationship between the race/ethnicity group of interest and the GVF parameters to use in standard error calculations.

Table 2. Estimation Groups of Interest and Generalized Variance Parameters	
Race/ethnicity group of interest	GVF parameters to use in standard error calculations
Total population	Total or White
White alone, White AOIC, or White non-Hispanic population	Total or White
Black alone, Black AOIC, or Black non-Hispanic population	Black
Asian alone, Asian AOIC, or Asian non-Hispanic population	Asian, AIAN, NHOPI
AIAN alone, AIAN AOIC, or AIAN non-Hispanic population	
NHOPI alone, NHOPI AOIC, or NHOPI non-Hispanic population	
Populations from other race groups	Asian, AIAN, NHOPI
Hispanic population	Hispanic
Two or more races – employment/unemployment and educational attainment characteristics	Black
Two or more races – all other characteristics	Asian, AIAN, NHOPI

- Notes: (1) AIAN is American Indian and Alaska Native and NHOPI is Native Hawaiian and Other Pacific Islander.
- (2) AOIC is an abbreviation for alone or in combination. The AOIC population for a race group of interest includes people reporting only the race group of interest (alone) and people reporting multiple race categories including the race group of interest (in combination).
- (3) Hispanics may be any race.
- (4) Two or more races refers to the group of cases self-classified as having two or more races.

Standard Errors of Estimated Numbers. The approximate standard error, s_x , of an estimated number from this microdata file can be obtained by using the formula:

$$s_x = \sqrt{ax^2 + bx} \quad (1)$$

Here x is the size of the estimate, and a and b are the parameters in Table 4 or 5 associated with the particular type of characteristic. When calculating standard errors from cross-tabulations involving different characteristics, use the set of parameters for the characteristic that will give the largest standard error.

Illustration 1

Suppose there were 3,362,000 unemployed men (ages 16 and up) in the civilian labor force. Use the appropriate parameters from Table 4 and Formula (1) to get

Illustration 1	
Number of unemployed males in the civilian labor force (x)	3,362,000
a-parameter (a)	-0.000031
b-parameter (b)	2,947
Standard error	98,000
90-percent confidence interval	3,201,000 to 3,523,000

The standard error is calculated as

$$s_x = \sqrt{-0.000031 \times 3,362,000^2 + 2,947 \times 3,362,000} = 98,000$$

The 90-percent confidence interval is calculated as $3,362,000 \pm 1.645 \times 98,000$.

A conclusion that the average estimate derived from all possible samples lies within a range computed in this way would be correct for roughly 90 percent of all possible samples.

Standard Errors of Estimated School Enrollment Numbers. The approximate standard error, s_x , of an estimated school enrollment number from this microdata file can be obtained by using the formula:

$$s_x = \sqrt{-\left(\frac{b}{T}\right)x^2 + bx} \quad (2)$$

Here x is the size of the estimate, T is the population total in Table 3 for the total number of persons in a specific age group, and b is the parameter in Table 5 associated with the particular type of characteristic. If Table 3 does not contain the age group of interest, use the smallest age group available in the table that does contain the age group of interest. When calculating standard errors for numbers from cross-tabulations involving different characteristics, use the set of parameters for the characteristic that will give the largest standard error.

Table 3. Population Totals for School Enrollment Age Groups: October 2017					
Age Group	Total	White	Black	Asian, AIAN, NHOPI	Hispanic
3+	310,070,226	237,735,230	40,971,396	31,363,600	58,243,443
3-4	9,165,161	5,716,228	1,835,784	1,613,149	5,168,456
3-6	18,350,106	11,491,170	3,670,481	3,188,455	10,425,046
3-17	71,067,484	50,774,462	11,112,141	9,180,881	20,443,173
3-24	92,382,926	66,410,796	14,324,312	11,647,818	25,131,352
5-24	83,217,765	60,694,568	12,488,528	10,034,669	19,962,896
6-13	34,048,740	23,808,268	5,557,731	4,682,741	11,409,765
14 -17	24,991,619	18,388,002	3,718,626	2,884,991	10,018,127
15+	260,368,881	202,323,543	32,977,628	25,067,710	47,818,397
15 -17	21,366,139	15,362,775	3,118,373	2,884,991	10,018,127
15-19	21,366,139	15,362,775	3,118,373	2,884,991	10,018,127
15-24	42,681,581	30,999,109	6,330,544	5,351,928	14,706,306
16-17	16,763,274	12,337,177	2,510,428	1,915,669	3,864,952
16-24	38,078,716	27,973,511	5,722,599	4,382,606	8,553,131
18-19	16,763,274	12,337,177	2,510,428	1,915,669	3,864,952
18-24	38,078,716	27,973,511	5,722,599	4,382,606	8,553,131
20-21	21,315,442	15,636,334	3,212,171	2,466,937	4,688,179
20-24	21,315,442	15,636,334	3,212,171	2,466,937	4,688,179
22-24	21,315,442	15,636,334	3,212,171	2,466,937	4,688,179
25+	217,687,300	171,324,434	26,647,084	19,715,782	33,112,091
25-29	22,731,454	16,612,710	3,459,013	2,659,731	4,651,311
25-34	44,090,874	32,585,700	6,355,648	5,149,526	9,024,526
30-34	21,359,420	15,972,990	2,896,635	2,489,795	4,373,215
35+	173,596,426	138,738,734	20,291,436	14,566,256	24,087,565

Notes: (1) AIAN is American Indian and Alaska Native and NHOPI is Native Hawaiian and Other Pacific Islander.

(2) White, Black, and Asian, AIAN, NHOPI totals include Hispanics.

(3) Hispanics may be any race.

Illustration 2

Suppose there were 4,319,000 three- and four-year-olds enrolled in school and 9,165,161 total children in that age group. Use the appropriate b parameter from Table 5 and Formula (2) to get

Illustration 2	
Number of three and four year olds enrolled in school (x)	4,319,000
Total (T)	9,165,161
b parameter (b)	2,912
Standard error	82,000
90-percent confidence interval	4,184,000 to 4,454,000

The standard error is calculated as

$$s_x = \sqrt{-\left(\frac{2,912}{9,165,161}\right) \times 4,319,000^2 + 2,912 \times 4,319,000} = 82,000$$

The 90-percent confidence interval is calculated as $4,319,000 \pm 1.645 \times 82,000$.

A conclusion that the average estimate derived from all possible samples lies within a range computed in this way would be correct for roughly 90 percent of all possible samples.

Standard Errors of Estimated Percentages. The reliability of an estimated percentage, computed using sample data for both numerator and denominator, depends on both the size of the percentage and its base. Estimated percentages are relatively more reliable than the corresponding estimates of the numerators of the percentages, particularly if the percentages are 50 percent or more. When the numerator and denominator of the percentage are in different categories, use the parameter from Table 4 or 5 as indicated by the numerator.

The approximate standard error, $s_{y,p}$, of an estimated percentage can be obtained by using the formula:

$$s_{y,p} = \sqrt{\frac{b}{y} p(100 - p)} \tag{3}$$

Here y is the total number of people, families, households, or unrelated individuals in the base or denominator of the percentage, p is the percentage $100 \times x/y$ ($0 \leq p \leq 100$), and b is the parameter in Table 4 or 5 associated with the characteristic in the numerator of the percentage.

Illustration 3

Suppose there were 16,533,000 people aged 18 to 21, and 50.3 percent were enrolled in college. Use the appropriate parameter from Table 5 and Formula (3) to get

Illustration 3	
Percentage of people aged 18-21 enrolled in college (<i>p</i>)	50.3
Base (<i>y</i>)	16,533,000
b-parameter (<i>b</i>)	2,530
Standard error	0.62
90-percent confidence interval	49.3 to 51.3

The standard error is calculated as

$$s_{y,p} = \sqrt{\frac{2,530}{16,533,000} \times 50.3 \times (100.0 - 50.3)} = 0.62$$

The 90-percent confidence interval for the estimated percentage of people aged 18 to 21 enrolled in college is from 49.3 to 51.3 percent (i.e., $50.3 \pm 1.645 \times 0.62$).

Standard Errors of Estimated Differences. The standard error of the difference between two sample estimates is approximately equal to

$$s_{x_1-x_2} = \sqrt{s_{x_1}^2 + s_{x_2}^2} \tag{4}$$

where s_{x_1} and s_{x_2} are the standard errors of the estimates, x_1 and x_2 . The estimates can be numbers, percentages, ratios, etc. This will result in accurate estimates of the standard error of the same characteristic in two different areas or for the difference between separate and uncorrelated characteristics in the same area. However, if there is a high positive (negative) correlation between the two characteristics, the formula will overestimate (underestimate) the true standard error.

Illustration 3

Suppose that of the 7,248,000 employed men between 20-24 years of age, 30.1 percent were part-time workers, and of the 6,932,000 employed women between 20-24 years of age, 39.4 percent were part-time workers. Use the appropriate parameters from Table 5 and Formulas (3) and (4) to get

Illustration 3			
	Men (x_1)	Women (x_2)	Difference
Percentage working part-time (p)	30.1	39.4	9.3
Base (y)	7,248,000	6,932,000	-
b-parameter (b)	2,947	2,788	-
Standard error	0.92	0.98	1.34
90-percent confidence interval	28.6 to 31.6	37.8 to 41.0	7.1 to 11.5

The standard error of the difference is calculated as

$$s_{x_1-x_2} = \sqrt{0.92^2 + 0.98^2} = 1.34$$

The 90-percent confidence interval around the difference is calculated as $9.3 \pm 1.645 \times 1.34$. Since this interval does not include zero, we can conclude with 90-percent confidence that the percentage of part-time women workers between 20-24 years of age is greater than the percentage of part-time men workers between 20-24 years of age.

Standard Errors of Quarterly or Yearly Averages. For information on calculating standard errors for labor force data from the CPS which involve quarterly or yearly averages, please see the “Explanatory Notes and Estimates of Error: Household Data” section in *Employment and Earnings*, a monthly report published by the U.S. Bureau of Labor Statistics.

Year-to-Year Factors.

In past years, the Census Bureau published a table of year factors for the School Enrollment Supplement in the Source and Accuracy Statement. User demand for these factors has diminished with the introduction of replicate weights. Data users producing estimates from prior years should consult the Source and Accuracy Statements covering the years of their analysis to estimate standard errors.

Technical Assistance. If you require assistance or additional information, please contact the Demographic Statistical Methods Division via e-mail at dsmd.source.and.accuracy@census.gov.

Table 4. Parameters for Computation of Standard Errors for Labor Force Characteristics: October 2017

Characteristic	<i>a</i>	<i>b</i>
Total or White		
<i>Civilian labor force, employed</i>	-0.000013	2,481
<i>Unemployed</i>	-0.000017	3,244
<i>Not in labor force</i>	-0.000013	2,432
<i>Civilian labor force, employed, not in labor force, and unemployed</i>		
Men	-0.000031	2,947
Women	-0.000028	2,788
Both sexes, 16 to 19 years	-0.000261	3,244
Black		
<i>Civilian labor force, employed, not in labor force, and unemployed</i>		
Total	-0.000117	3,601
Men	-0.000249	3,465
Women	-0.000191	3,191
Both sexes, 16 to 19 years	-0.001425	3,601
Asian, American Indian and Alaska Native, Native Hawaiian and Other Pacific Islander		
<i>Civilian labor force, employed, not in labor force, and unemployed</i>		
Total	-0.000245	3,311
Men	-0.000537	3,397
Women	-0.000399	2,874
Both sexes, 16 to 19 years	-0.004078	3,311
Hispanic, may be of any race		
<i>Civilian labor force, employed, not in labor force, and unemployed</i>		
Total	-0.000087	3,316
Men	-0.000172	3,276
Women	-0.000158	3,001
Both sexes, 16 to 19 years	-0.000909	3,316

- Notes: (1) These parameters are to be applied to basic CPS monthly labor force estimates.
(2) The Total or White, Black, and Asian, AIAN, NHOPI parameters are to be used for both alone and in combination race group estimates.
(3) For nonmetropolitan characteristics, multiply the a- and b-parameters by 1.5. If the characteristic of interest is total state population, not subtotaled by race or ethnicity, the a- and b-parameters are zero.
(4) For foreign-born and noncitizen characteristics for Total and White, the a- and b-parameters should be multiplied by 1.3. No adjustment is necessary for foreign-born and noncitizen characteristics for Black, Hispanic, and Asian, AIAN, NHOPI parameters.

- (5) For the groups self-classified as having two or more races, use the Asian, AIAN, NHOPI parameters for all employment characteristics.

Table 5. Parameters for Computation of Standard Errors for School Enrollment Characteristics: October 2017				
Characteristics	<i>b</i>			
	Total or White	Black	Asian, AIAN, NHOPI	Hispanic
PEOPLE				
<i>Persons enrolled in school:</i>				
Total.....	2,530	2,861	2,861	3,258
Children 13 and under.....	2,912	3,295	3,295	3,750
<i>Marital status, household and family characteristics, health insurance</i>				
Some household members.....	5,564	7,993	7,993	13,471
All household members.....	6,761	11,788	11,788	19,866
FAMILIES, HOUSEHOLDS, OR UNRELATED INDIVIDUALS				
<i>Income, earnings.....</i>	2,393	2,613	2,613	4,403
<i>Marital status, household and family characteristics, educational attainment, population by age/sex.....</i>	2,208	1,998	1,998	3,367

- Notes: (1) These parameters are to be applied to the October 2017 School Enrollment Supplement data.
(2) AIAN is American Indian and Alaska Native and NHOPI is Native Hawaiian and Other Pacific Islander.
(3) Hispanics may be any race. For a more detailed discussion on the use of parameters for race and ethnicity, please see the “Generalized Variance Parameters” section.
(4) The Total or White, Black, and Asian, AIAN, NHOPI parameters are to be used for both alone and in combination race group estimates.
(5) For nonmetropolitan characteristics, multiply the a- and b-parameters by 1.5. If the characteristic of interest is total state population, not subtotaled by race or ethnicity, the a- and b-parameters are zero.
(6) For foreign-born and noncitizen characteristics for Total and White, the a- and b-parameters should be multiplied by 1.3. No adjustment is necessary for foreign-born and noncitizen characteristics for Black, Asian, AIAN, NHOPI, and Hispanic parameters.
(7) For the group self-classified as having two or more races, use the Asian, AIAN, NHOPI parameters for all characteristics except employment, unemployment, and educational attainment, in which case use Black parameters.

Table 6. Factors for Region Parameters: October 2017

Geography	Factor
U.S. Totals	1.00
Regions:	
Northeast	1.08
Midwest	1.09
South	1.11
West	1.03

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